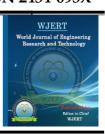


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# NUMBER OF ZEROS OF A CLASS OF RANDOM ALGEBRAIC POLYNOMIAL

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# **ABSTRACT**

The expected number of crossings of a Random algebraic polynomial  $f(n)\to\infty$ , crosses the line y=mx, when m is any real value and  $(m^2/n)\to 0$  as  $n\to\infty$  reduces to only one. We know the expected number of times that a polynomial of degree n with independent normally distributed random real coefficients asymptotically crosses the line y=mx, when m is any real value and  $(m^2/n)\to 0$  as  $n\to\infty$ . Many authors Kac, [4] Farahmad, [3] Nayak. [7] Rice [9] have investigated the plane symmetric solutions of number of crossings of different polynomials and equations in general relativity. Here, we study the expected number of crossings of a Random algebraic polynomial f (n)

 $\to\infty$ , crosses the line y= mx, when m is any real value and (m²/n)  $\to 0$  as n $\to\infty$  reduces to only one. This theory agrees with the present observational facts pertaining to general relativity. The present paper shows that the expected number of crossings of a Random algebraic polynomial for m >exp(f(n)), where f is any function of n such that f(n) $\to\infty$ , this expected number of crossings reduces to onlyone.

**KEYWORDS AND PHRASES:** Independent identically distributed random variables, random algebraic polynomial, random algebraic equation, real roots.

#### 1. INTRODUCTION

Consider the algebraic polynomial

$$P(x) = a_0 x_0 + a_1 x_1 + \dots + a_n x_n \tag{1.1}$$

where a<sub>0</sub>,a<sub>1</sub>,a<sub>2</sub>,....a<sub>n</sub>is a sequence of independent, normally distributed random variables with mathematical exception  $\mu$  and variance unity. The set of equations y=P(x) represents a family of curves in the xy-plane, Kac, [3] and Cramer shows that for  $\mu$ =0 the number of times that this family crosses the line x-axis, on a average is  $(2/\pi) \log n$ . Mishra<sup>[5]</sup> obtained the same average number of crossings when they considered the case of coefficients belongings to the domain of attraction of the normal law with mean  $\mu$ =0. They also showed that when  $\mu=0$  the number of crossings reduces by half. Mohanty<sup>[7]</sup> and Ibragimov<sup>[2]</sup> studied the number of times that this family of curves crosses the level K=0( $\sqrt{n}$ ) (crosses with line y=K) for μ=0 and showed that these numbers decreases as K increases. He also showed that even in this case for  $\mu$ the number of crossings reduces by half. Denote by  $N_m$  (a,b)= N (a,b) the number of times that this family crosses the line y=mx where m is any constant independent of x and let EN (a,b) be its expectation. For  $m = 0(\sqrt{n})$  an asymptotic value for  $EN(-\infty,\infty)$  was obtained by Mohanty, [7] and Pratihari [8] with which the reader will be assumed to be familiar. As noted in the latter there is a sizeable number of crossings even when the line tends to be perpendicular to the x axis, that is for m (=0 ( $\sqrt{n}$ ) $\rightarrow \infty$  as n $\rightarrow \infty$ . In this work we study the case when m is very large compared with n, and show that the number of crossings of this family of curves with such a line reduces to one. We prove.

**Theorem-** If the coefficients of P(x) in (1.1) are independent normally distributed random variables with mean zero and variance unity, then for any constant m such that  $|m| > \exp(f(n))$ , where f is any function of n such that f(n) tends to infinity as n tends to infinity, the mathematical expectation of the number of real roots of the equation y=P(x)=mx is asymptotic to one.

2. Proof of the theorem:-First we find a lower estimates for EN  $(-\infty,\infty)$ . Let m>exp (n,f), then since for |x| < 1the polynomial P(x) is convergent, with probability one, for x=1/2, say and n sufficiently large.

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P(x) = mx < P(1/2) - (1/2) \exp(nf) < 0
and also for x = -1/2
P(x) = -mx > P(-1/2) + (1/2) \exp(nf) > 0
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Therefore, by the intermediate value theorem, there exists at least one real root for the function P(x)-mx in the interval  $(-1/2, \frac{1}{2})$ . Similarly, if m<-exp (nf) we can show that the function P(x)-mx takes on the opposite sign at x=1/2 and x=-1/2, therefore, there exists at least one real root. Hence EN  $(-\infty, \infty) \ge 1$  and we only have to show that the upper limit

is one as well. We also note that both aj and -aj (i=0,1,2,....n-1) have standard normal distribution hence changing x to -x leaves the distribution of the coefficients invariant, thus EN ( $-\infty$ ,0)=EN ( $-\infty$ ,0). So we only have to consider the interval ( $0,\infty$ ). In by using the expected number of level crossings, the Kac-Rice formula<sup>[4]</sup> for the equation P(x)- mx=0 is found.

$$EN(a,b) = \int_{a}^{b} \left[ (\Delta^{1/2} / \pi A) \exp\{(-Am^{2} + 2m^{2}Cx - Bm^{2}x^{2}) / 2D \right]$$

$$+ \left( \sqrt{2/n} \right) m(Dx - A) |A^{-a/2}| \exp(m_{2}x^{2} / 2A) erf$$

$$\left\{ m(Dx - A) (2A\Delta)^{-1/2} \right\} dx$$

$$= \int_{a}^{b} I_{1}(x) dx + \int_{a}^{b} I_{2}(x) dx \qquad \text{say} \qquad (2.1)$$

Where

$$A = \sum_{i=0}^{n-1} {n-1 \choose i} x^{2i} = (1+x^2)^{n-1}$$

$$(2.2)$$
Since  $\frac{d}{dt} (1+t)^{n-1} = \frac{d}{dt} \sum_{i=0}^{n-1} {n-1 \choose i} t^i = \sum_{i=0}^{n-1} {n-1 \choose i} t^{i-1}$ 

$$\Rightarrow t(n-1)(1+t)^{n-2} = \sum_{i=0}^{n-1} i {n-1 \choose i} t^i$$
Similarly  $\frac{d^2}{dt^2} (1+t)^{n-1} = \sum_{i=0}^{n-1} i (i-1) {n-1 \choose i} t^{i-2}$ 

$$\Rightarrow t^2 (n-1)(n-2)(1+t)^{n-3} = \sum_{i=0}^{n-1} i (i-1) {n-1 \choose i} t^{i}$$

$$\Rightarrow t^2 (n-1)(n-2)(1+t)^{n-3} + t(n-1)(1+t)^{n-2} = \sum_{i=0}^{n-1} i^2 {n-1 \choose i} t^i$$

$$Again \frac{d^3}{dt^3} (1+t)^{n-1} = \sum_{i=0}^{n-1} i (i-1)(i-2) {n-1 \choose i} t^{i-3}$$

$$\Rightarrow t^3 (n-1)(n-2)(n-3)(1+t)^{n-4} = \sum_{i=0}^{n-1} i^3 - 3i^2 + 2i {n-1 \choose i} t^i$$

$$\Rightarrow t^3 (n-1)(n-2)(n-3)(1+t)^{n-4} + 3t^2 (n-1)(n-2)(1+t)^{n-3}$$

$$+ t(n-1)(1+t)^{n-2} = \sum_{i=0}^{n-1} i^3 {n-1 \choose i} t^i$$

$$Again \frac{d^4}{dt^4} (1+t)^{n-1} = \sum_{i=0}^{n-1} i (i-1)(i-2)(i-3) {n-1 \choose i} t^{i-4}$$

$$\Rightarrow t^4 (n-1)(n-2)(n-3)(n-4)(1+t)^{n-5} = \sum_{i=0}^{n-1} (i^4 - 6i^3 + 11i^2 - 6i) {n-1 \choose i} t^i$$

$$\Rightarrow t^4 (n-1)(n-2)(n-3)(n-4)(1+t)^{n-5} + 6t^3 (n-1)(n-2)(n-3)(1+t)^{n-4}$$

$$+7t^2 (n-1)(n-2)(1+t)^{n-3} + t(n-1)(1+t)^{n-2} = \sum_{i=0}^{n-1} i^4 {n-1 \choose i} t^i$$
Then  $B = \sum_{i=0}^{n-1} i^2 {n-1 \choose i} x^{2i-2}$ 

$$= x^{-2} \sum_{i=0}^{n-1} i^2 {n-1 \choose i} x^{2i}$$

$$= x^{-2} (x^4 (n-1)(n-2)(1+x^2)^{n-3} + x^2 (n-1)(1+x^2)^{n-2}$$

$$= (n-1)(1+x^2)^{n-3} (x^2 (n-2)+1+x^2)$$

$$= (n-1)(1+x^2)^{n-3} (x^2 (n-2)+1+x^2)$$

$$= (n-1)(1+x^2)^{n-3} (nx^2 - x^2 + 1)$$

$$C = \sum_{i=0}^{n-1} i^2 (i-1)^2 \binom{n-1}{i} x^{2i-4}$$

$$= x^{-4} \sum_{i=0}^{n-1} (i^4 + i^2 - 2i^3) \binom{n-1}{i} x^{2i}$$

$$= x^{-4} (x^8 (n-1)(n-2)(n-3)(n-4)(1+x^2)^{n-5} + 4x^6 (n-1)(n-2)(n-3)(1+x^2)^{n-4} + 2x^4 (n-1)(n-2)(1+x^2)^{n-3})$$

$$= (n-1)(n-2) (1+x^2)^{n-5} \{x^4 (n-3)(n-4) + 4x^2 (n-3)(1+x^2) + 2(1+x^2)^2\}$$

$$= (n-1)(n-2) (1+x^2)^{n-5} \{n^2 x^4 - 3nx^4 + 2x^4 + 4nx^2 - 8x^2 + 2\}$$

$$= (n-1)(n-2) (1+x^2)^{n-5} \{(n-1)(n-2)x^4 + 4(n-2)x^2 + 2\}$$

$$= x^{-1} (n-1)(n-2) (1+x^2)^{n-5} \{(n-1)(n-2)x^4 + 4(n-2)x^2 + 2\}$$

$$= x^{-1} \sum_{i=0}^{n-1} i \binom{n-i}{i} x^{2i-1}$$

$$= x^{-1} \sum_{i=0}^{n-1} i \binom{n-i}{i} x^{2i}$$

$$= x^{-1} (x^2 (n-1)(1+x^2)^{n-2}) = x(n-1)(1+x^2)^{n-2}$$

$$= x^{-2} (\sum_{i=0}^{n-1} i^2 \binom{n-1}{i} x^{2i-2}$$

$$= x^{-2} (\sum_{i=0}^{n-1} i^2 \binom{n-1}{i} x^{2i} - \sum_{i=0}^{n-1} i \binom{n-1}{i} x^{2i})$$

$$= x^{-2} (x^4 (n-1)(n-2)(1+x^2)^{n-3})$$

$$= x^2 (n-1)(n-2)(1+x^2)^{n-3}$$

$$= x^2 (n-1)(n-2)(1+x^2)^{n-3}$$

$$= x^{-3} (x^6 (n-1)(n-2)(n-3)(1+x^2)^{n-4} + x^4 (n-1)(n-2)(1+x^2)^{n-3})$$

$$= (x^3 (n-1)(n-2)(n-3)(1+x^2)^{n-4} + 2x(n-1)(n-2)(1+x^2)^{n-3})$$

$$= (n-1)(n-2)(1+x^2)^{n-4} (nx^3 - x^3 + 2x)$$
(2.7)

and

$$erf(x) = \int_{0}^{x} \exp(-y^{2}) dy$$
 (2.8)

First we show  $\int_{0}^{1} I_1(x) dx$  tends to zero as  $n \to \infty$ . Let a be constant independent of x in the interval (0,1). For  $0 \le x \le 1 - n^{-\alpha}$  and sufficiently large n we have

$$D = \{ (n-1)x^{2n+1} - nx^{2n-1} + x \} (1-x^2)^{-2}$$
  
=  $x(1-x^{2n})(1-x^2)^{-2} + 0\{n^{1+\alpha} \exp(-2n^{1-\alpha})\}....(2.9)$ 

and

$$erf = \int_{0}^{x} exp(-y^{2}) dy$$

First we show that  $\int_0^1 \mathbf{I}_1(\mathbf{x}) d\mathbf{x}$  tends to zero as  $n \to \infty$ . Let a be a constant independent of x in the interval (0,1). For  $0 \le x \le 1 - n^{-a}$  and all sufficiently large nwe have

$$D = \{(n-1)x^{2n+1} - nx^{2n-1} + x\}(1-x^2)^{-2}$$
  
=  $x(1-x^{2n})(1-x^2)^{-2} + 0\{n^{1+\alpha}\exp(-2n^{1-\alpha})\}....(2.9)$ 

and

$$B = (1+x^2)(1-x^{2n})(1-x^{2n})(1-x^{2n})^{-3} + 0\{n^{2+a}\exp(-2n^{1-a})\}$$
(2.10)

From (2.8) and (2.9) we can obtain

$$\Delta = (1 - x^{2n})(1 - x^2)^{-4} + 0\{n^{2+2n} \exp(-2n^{1-a})\}$$
(2.11)

Now we choose  $a=1-\{\log\log{(n)}^{10}\}/\log{n}$ . Then since, for all sufficiently large

n, 
$$n^{2+a} \exp(-2n^{1-a}) = n^{2+2n} \exp(-2\log(n)^{10}) = n^{-18+a} \to 0$$
.

All the error terms that appear in the formulas (2.9) to (2.11) will tend to zero.

Hence from (2.1), (2.9), (2.10), (2.11) and since for all x

$$(1-x^2)^{2/2} - x^2(1-x^2) + x^2(1+x^2)/2 > 1/5$$

we have

$$I_{1}(x) = \int_{0}^{1-n^{-a}} \left( \Delta^{1/2} / wA \right) \exp\left[ -m^{2} \left\{ (1-x^{2})/(1-x^{2n}) \right\} \right]$$

$$\times \left\{ (1-x^{2})^{2} / 2 - x2(1-x^{2}) \right\}$$

$$+ x^{2} (1+x^{2})/2 \left\{ (1+0 \left\{ n^{2+n} \exp\left(-2n^{1-a}\right) \right\} dx \right\}$$

$$\leq (1/\pi) \int_{0}^{1-n^{-a}} (1-x^{2})^{-1} \exp\left\{ -m^{2} (1-x^{2})/5 \right\} dx$$

$$\leq (a/2\pi) \exp\left\{ \log \log n \exp\left(-(m^{2}/2n) \log (n)^{10} \right\} \dots$$
(2.12)

Now we note that since m>exp (n/logn) the term  $m^2/n$  tends to infinity much faster than log log n as  $n\to\infty$ , hence from (2.12) we can obtain

$$\int_{0}^{1-n^{-a}} I_{1}(x)dx \to 0 \text{ as } n \to \infty.$$
 (2.13)

To show that  $\int_{0}^{1-n^{-a}} I_1(x)dx \to 0 \text{ as } n \to \infty.$  we first prove that  $(A-2Dx+Bx^2)/\Delta$  is

positive for  $1 - n^{-a} \le x \le 1$ . For all sufficiently large n from (2.2) we have

$$A - 2Dx + Bx^{2} \ge Bx^{2} - 2Dx$$

$$\ge \left\{ n^{2}x^{2n}(1 - x^{2})2 - 2nx^{2n+2}(1 - x^{2}) \right\}$$

$$+ x^{2}(1 + x^{2})(1 - x^{2n}) \} (1 - x^{2}) - 3 - n(n+1)$$

$$\ge n^{3} \left\{ \log(n)^{10} \right\}^{-3} - 2n^{2} > n^{2}$$
(2.14)

since  $(1-x^2)^3 < (2n^{-a})^3$  and  $2nx^{2n+2}(1-x^2) \to as \ n \to \infty$ .

Hence from (2.14) and since  $\Delta < n^4$  we have

$$(A - 2Dx + Bx^{2})/\Delta > n^{-2}$$
 (2.15)

So from (2.15) and since from  $(\Delta^2 / a) < (2n-1)^{1/2} (1-x)^{-1/2}$ ,

we have 
$$\int_{0}^{1-n^{-a}} I_{1}(x)dx < (2n-1)^{1/2} \exp(-m^{2}/n^{2}) \int_{0}^{1-n^{-a}} (1-x)^{-1/2} dx$$

$$< 3n^{(1-a)}/2 \exp(-m^{2}/n^{2})$$
(2.16)

Which tends to zero as  $n \rightarrow \infty$ 

In order to find  $\int_0^{1-n^{-\alpha}} I_1(x) dx$  we let y = 1/x, and divide the interval  $0 \le y \le 1$  into three subintervals (0,b), (b,1-(1/nd)) and (1-(1/nd),1) where  $d=\{(3/8) \log \log(n)^{1/3} \text{ and } b=(m^{-2}n \text{ d log } n)^{1/(4n-8)}$ . We show that in these three subintervals

$$\int_{0}^{1-n^{-a}} I_{1}(x)dx = \int_{0}^{1} y^{-2} I_{1}(1/y) dy = \log \{(1+b)/(1-b)\} \to 0 \text{ as } n \to \infty.$$
 (2.17)

For  $b \le y \le 1 - 1/nd$  from (2.2) we have

$$\{(A-2D)/D\} + B/D^2 = 1 + \sum_{i=2}^{n} y^{-2i} + \sum_{i=2}^{n} y^{-2i} (i^2 - 2i) > n^3/4$$
 (2.18)

and

$$\Delta = \{1 - h^{2}(y)\} (1 - y^{2n}) 2 / y^{4n-8} (1 - y^{2}) 4 \le n^{2} / b^{4n-8} (1 - y^{2})^{2}$$

$$< 3n^{4} d^{2} / b^{4n-8}$$
(2.19)

To prove that

$$\int_{0}^{\infty} I_{1}(x) dx \to 0 \text{ as } n \to \infty, \text{ from (2.2) for y = 1/x we have }.$$

$$(Dx - A) / A^{1/2} \Delta^{1/2} = D^{2n-1} (1 - y^{2})^{3/2} \{ (n - 1 - ny - y^{2n-1}) (1 - y^{2})^{2} \}$$

$$- y^{2n-4} \} \{ 1 - h^{2}(y) \}^{-1/2} (1 - y^{2n})^{-1/2}$$
(2.20)

And to find 
$$\int_{a}^{b} I_{2}(x) dx$$

Let 
$$x^2/A = y^{2n-4}(1-y^2)/(1-y^{2n})$$
 (2.21)

First we let

$$0 \le y \le (mn^2)^{-1/(2n-1)}$$
 then  $y \to 0$  as  $n \to \infty$ , from (2.21)

we have 
$$(Dx - A)/A^{1/2}\Delta^{1/2} < 2ny^{2n-1}$$
 (2.22)  

$$\int_{0}^{1} y^{-2} \mathbf{I}_{2} (1/y) dy < \int_{\mathbf{m}/2n^{2}}^{\infty} \exp(-\mathbf{u}^{2}/2) d\mathbf{u}$$

$$\leq (m/\sqrt{n-m^{2}/2n^{2}}) \exp(-m^{2}/4n^{4}) \to 0 \text{ as } n \to \infty.$$
 (2.23)

Hence from (2.22) and (2.23) we have

$$\int_{1}^{\infty} I_2(x) dx \to 0 \text{ as } n \to \infty.$$
 (2.24)

Finally from (2.8), (2.13), (2.16), (2.18) and (2.24) we obtain

$$_{EN}(0,\infty) \leq 1/2$$

and since  $EN(-\infty, \infty) = 2EN(0, \infty)$  we have proof of the theorem.

## **RESULT**

The asymptotic number of crossings of the polynomial P(x) with line y=mx decreases as  $m=0(\sqrt{n})$  increases. In this paper we proved that when  $|m| \ge \exp(nf)|$  the number of crossings reduces to one. The behaviour of the number of crossings between these two lines is not known. A subsequent studycould consider the case when  $(m^2/n)$  tends to any non zero constant as n tends to infinity and as a guessed target  $EN(-\infty,\infty) \sim (1/\pi)\log n$ , which is half the number of crossings when m=0 seems reasonable. (Knowing a rough value for  $EN(-\infty,\infty)$  is useful in order to sufficient upper and lower bounds for  $EN(-\infty,\infty)$  leading to an asymptotic formula). Indeed, the behaviour of  $EN(-\infty,\infty)$  for other values of m is also interesting, but it will involve more analysis especially for the  $\int_{-\infty}^{\infty} I_2(x) dx$  part of  $EN(-\infty,\infty)$ . The result of the theorem shows that the expected number of crossings of a Random algebraic polynomial form  $>\exp(f(n))$ , where f is any function of n such that  $f(n)\to\infty$ , this expected number of crossings reduces to only one.

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