



NUMBER OF ZEROS OF A CLASS OF RANDOM ALGEBRAIC POLYNOMIAL

Sima Rout* and P. K. Mishra

Department of Mathematics Odisha University of Technology and Research,
Technocampusmahalaxmi Vihar, Nakagate, Khandagiri, Bhubaneswar, Odisha.

Article Received on 15/06/2024

Article Revised on 05/07/2024

Article Accepted on 25/07/2024



*Corresponding Author

Sima Rout

Department of Mathematics
Odisha University of
Technology and Research,
Technocampusmahalaxmi
Vihar, Nakagate,
Khandagiri, Bhubaneswar,
Odisha.

ABSTRACT

The expected number of crossings of a Random algebraic polynomial $f(n) \rightarrow \infty$, crosses the line $y = mx$, when m is any real value and $(m^2/n) \rightarrow 0$ as $n \rightarrow \infty$ reduces to only one. We know the expected number of times that a polynomial of degree n with independent normally distributed random real coefficients asymptotically crosses the line $y = m x$, when m is any real value and $(m^2/n) \rightarrow 0$ as $n \rightarrow \infty$. Many authors Kac,^[4] Farahmad,^[3] Nayak.^[7] Rice^[9] have investigated the plane symmetric solutions of number of crossings of different polynomials and equations in general relativity. Here, we study the expected number of crossings of a Random algebraic polynomial $f(n)$

$\rightarrow \infty$, crosses the line $y = mx$, when m is any real value and $(m^2/n) \rightarrow 0$ as $n \rightarrow \infty$ reduces to only one. This theory agrees with the present observational facts pertaining to general relativity. The present paper shows that the expected number of crossings of a Random algebraic polynomial for $m > \exp(f(n))$, where f is any function of n such that $f(n) \rightarrow \infty$, this expected number of crossings reduces to only one.

KEYWORDS AND PHRASES: Independent identically distributed random variables, random algebraic polynomial, random algebraic equation, real roots.

1. INTRODUCTION

Consider the algebraic polynomial

$$P(x) = a_0x_0 + a_1x_1 + \dots + a_nx_n \quad (1.1)$$

where $a_0, a_1, a_2, \dots, a_n$ is a sequence of independent, normally distributed random variables with mathematical expectation μ and variance unity. The set of equations $y=P(x)$ represents a family of curves in the xy -plane, Kac,^[3] and Cramer^[1] shows that for $\mu=0$ the number of times that this family crosses the line x -axis, on an average is $(2/\pi) \log n$. Mishra^[5] obtained the same average number of crossings when they considered the case of coefficients belonging to the domain of attraction of the normal law with mean $\mu=0$. They also showed that when $\mu=0$ the number of crossings reduces by half. Mohanty^[7] and Ibragimov^[2] studied the number of times that this family of curves crosses the level $K=0(\sqrt{n})$ (crosses with line $y=K$) for $\mu=0$ and showed that these numbers decrease as K increases. He also showed that even in this case for μ the number of crossings reduces by half. Denote by $N_m(a,b) = N(a,b)$ the number of times that this family crosses the line $y=mx$ where m is any constant independent of x and let $EN(a,b)$ be its expectation. For $m = 0(\sqrt{n})$ an asymptotic value for $EN(-\infty, \infty)$ was obtained by Mohanty,^[7] and Pratihari^[8] with which the reader will be assumed to be familiar. As noted in the latter there is a sizeable number of crossings even when the line tends to be perpendicular to the x axis, that is for $m (=0(\sqrt{n})) \rightarrow \infty$ as $n \rightarrow \infty$. In this work we study the case when m is very large compared with n , and show that the number of crossings of this family of curves with such a line reduces to one. We prove.

Theorem- If the coefficients of $P(x)$ in (1.1) are independent normally distributed random variables with mean zero and variance unity, then for any constant m such that $|m| > \exp(f(n))$, where f is any function of n such that $f(n)$ tends to infinity as n tends to infinity, the mathematical expectation of the number of real roots of the equation $y=P(x)=mx$ is asymptotic to one.

2. Proof of the theorem:-First we find a lower estimate for $EN(-\infty, \infty)$. Let $m > \exp(n, f)$, then since for $|x| < 1$ the polynomial $P(x)$ is convergent, with probability one, for $x=1/2$, say and n sufficiently large.

$$P(x) = mx < P(1/2) - (1/2) \exp(nf) < 0$$

and also for $x = -1/2$

$$P(x) = -mx > P(-1/2) + (1/2) \exp(nf) > 0$$

Therefore, by the intermediate value theorem, there exists at least one real root for the function $P(x)-mx$ in the interval $(-1/2, 1/2)$. Similarly, if $m < -\exp(nf)$ we can show that the function $P(x)-mx$ takes on the opposite sign at $x=1/2$ and $x=-1/2$, therefore, there exists at least one real root. Hence $EN(-\infty, \infty) \geq 1$ and we only have to show that the upper limit

is one as well. We also note that both a_j and $-a_j$ ($i=0,1,2,\dots,n-1$) have standard normal distribution hence changing x to $-x$ leaves the distribution of the coefficients invariant, thus $EN(-\infty,0)=EN(-\infty,0)$. So we only have to consider the interval $(0,\infty)$. In by using the expected number of level crossings, the Kac-Rice formula^[4] for the equation $P(x)-mx=0$ is found.

$$\begin{aligned}
 EN(a,b) &= \int_a^b \left[(\Delta^{1/2} / \pi A) \exp \{ (-Am^2 + 2m^2Cx - Bm^2x^2) / 2D \} \right. \\
 &\quad + \left. \left(\sqrt{2/\pi} \right) m(Dx - A) | A^{-a/2} \exp(m_2x^2 / 2A) \operatorname{erf} \right. \\
 &\quad \left. \left\{ m(Dx - A) / (2A\Delta)^{1/2} \right\} dx \right. \\
 &= \int_a^b I_1(x) dx + \int_a^b I_2(x) dx \quad \text{say} \quad (2.1)
 \end{aligned}$$

Where

$$A = \sum_{i=0}^{n-1} \binom{n-1}{i} x^{2i} = (1+x^2)^{n-1} \quad (2.2)$$

$$\text{Since } \frac{d}{dt} (1+t)^{n-1} = \frac{d}{dt} \sum_{i=0}^{n-1} \binom{n-1}{i} t^i = \sum_{i=0}^{n-1} \binom{n-1}{i} i t^{i-1}$$

$$\Rightarrow t(n-1)(1+t)^{n-2} = \sum_{i=0}^{n-1} i \binom{n-1}{i} t^i$$

$$\text{Similarly } \frac{d^2}{dt^2} (1+t)^{n-1} = \sum_{i=0}^{n-1} i(i-1) \binom{n-1}{i} t^{i-2}$$

$$\Rightarrow t^2(n-1)(n-2)(1+t)^{n-3} = \sum_{i=0}^{n-1} i(i-1) \binom{n-1}{i} t^i$$

$$\Rightarrow t^2(n-1)(n-2)(1+t)^{n-3} + t(n-1)(1+t)^{n-2} = \sum_{i=0}^{n-1} i^2 \binom{n-1}{i} t^i$$

$$\text{Again } \frac{d^3}{dt^3} (1+t)^{n-1} = \sum_{i=0}^{n-1} i(i-1)(i-2) \binom{n-1}{i} t^{i-3}$$

$$\Rightarrow t^3(n-1)(n-2)(n-3)(1+t)^{n-4} = \sum_{i=0}^{n-1} (i^3 - 3i^2 + 2i) \binom{n-1}{i} t^i$$

$$\Rightarrow t^3(n-1)(n-2)(n-3)(1+t)^{n-4} + 3t^2(n-1)(n-2)(1+t)^{n-3}$$

$$+ t(n-1)(1+t)^{n-2} = \sum_{i=0}^{n-1} i^3 \binom{n-1}{i} t^i$$

$$\text{Again } \frac{d^4}{dt^4} (1+t)^{n-1} = \sum_{i=0}^{n-1} i(i-1)(i-2)(i-3) \binom{n-1}{i} t^{i-4}$$

$$\Rightarrow t^4(n-1)(n-2)(n-3)(n-4)(1+t)^{n-5} = \sum_{i=0}^{n-1} (i^4 - 6i^3 + 11i^2 - 6i) \binom{n-1}{i} t^i$$

$$\Rightarrow t^4(n-1)(n-2)(n-3)(n-4)(1+t)^{n-5} + 6t^3(n-1)(n-2)(n-3)(1+t)^{n-4}$$

$$+ 7t^2(n-1)(n-2)(1+t)^{n-3} + t(n-1)(1+t)^{n-2} = \sum_{i=0}^{n-1} i^4 \binom{n-1}{i} t^i$$

$$\text{Then } B = \sum_{i=0}^{n-1} i^2 \binom{n-1}{i} x^{2i-2}$$

$$= x^{-2} \sum_{i=0}^{n-1} i^2 \binom{n-1}{i} x^{2i}$$

$$= x^{-2} (x^4(n-1)(n-2)(1+x^2)^{n-3} + x^2(n-1)(1+x^2)^{n-2})$$

$$= x^2(n-1)(n-2)(1+x^2)^{n-3} + (n-1)(1+x^2)^{n-2}$$

$$= (n-1)(1+x^2)^{n-3} (x^2(n-2) + 1+x^2)$$

(2.3)

$$= (n-1)(1+x^2)^{n-3} (nx^2 - x^2 + 1)$$

$$\begin{aligned}
 C &= \sum_{i=0}^{n-1} i^2 (i-1)^2 \binom{n-1}{i} x^{2i-4} \\
 &= x^{-4} \sum_{i=0}^{n-1} (i^4 + i^2 - 2i^3) \binom{n-1}{i} x^{2i} \\
 &= x^{-4} (x^8(n-1)(n-2)(n-3)(n-4)(1+x^2)^{n-5} + 4x^6(n-1)(n-2)(n-3)(1+x^2)^{n-4} \\
 &\quad + 2x^4(n-1)(n-2)(1+x^2)^{n-3}) \\
 &= (n-1)(n-2)(1+x^2)^{n-5} \{x^4(n-3)(n-4) + 4x^2(n-3)(1+x^2) + 2(1+x^2)^2\} \\
 &= (n-1)(n-2)(1+x^2)^{n-5} \{n^2x^4 - 3nx^4 + 2x^4 + 4nx^2 - 8x^2 + 2\} \\
 &= (n-1)(n-2)(1+x^2)^{n-5} \{(n-1)(n-2)x^4 + 4(n-2)x^2 + 2\} \tag{2.4}
 \end{aligned}$$

$$\begin{aligned}
 D &= \sum_{i=0}^{n-1} i \binom{n-1}{i} x^{2i-1} \\
 &= x^{-1} \sum_{i=0}^{n-1} i \binom{n-1}{i} x^{2i} \\
 &= x^{-1} (x^2(n-1)(1+x^2)^{n-2}) = x(n-1)(1+x^2)^{n-2} \tag{2.5}
 \end{aligned}$$

$$\begin{aligned}
 E &= \sum_{i=0}^{n-1} i(i-1) \binom{n-1}{i} x^{2i-2} \\
 &= x^{-2} (\sum_{i=0}^{n-1} i^2 \binom{n-1}{i} x^{2i} - \sum_{i=0}^{n-1} i \binom{n-1}{i} x^{2i}) \\
 &= x^{-2} (x^4(n-1)(n-2)(1+x^2)^{n-3}) \\
 &= x^2(n-1)(n-2)(1+x^2)^{n-3} \tag{2.6}
 \end{aligned}$$

$$\begin{aligned}
 F &= \sum_{i=0}^{n-1} i^2 (i-1) \binom{n-1}{i} x^{2i-3} = x^{-3} (\sum_{i=0}^{n-1} i^3 \binom{n-1}{i} x^{2i} - \sum_{i=0}^{n-1} i^2 \binom{n-1}{i} x^{2i}) \\
 &= x^{-3} (x^6(n-1)(n-2)(n-3)(1+x^2)^{n-4} + x^4(n-1)(n-2)(1+x^2)^{n-3}) \\
 &= (x^3(n-1)(n-2)(n-3)(1+x^2)^{n-4} + 2x(n-1)(n-2)(1+x^2)^{n-3}) \\
 &= (n-1)(n-2)(1+x^2)^{n-4} (nx^3 - x^3 + 2x) \tag{2.7}
 \end{aligned}$$

and

$$\operatorname{erf}(x) = \int_0^x \exp(-y^2) dy \tag{2.8}$$

First we show $\int_0^1 I_1(x) dx$ tends to zero as $n \rightarrow \infty$. Let a be constant independent of x in the interval (0,1). For $0 \leq x \leq 1 - n^{-\alpha}$ and sufficiently large n we have

$$\begin{aligned}
 D &= \{(n-1)x^{2n+1} - nx^{2n-1} + x\}(1-x^2)^{-2} \\
 &= x(1-x^{2n})(1-x^2)^{-2} + 0\{n^{1+\alpha} \exp(-2n^{1-\alpha})\} \dots \dots \dots (2.9)
 \end{aligned}$$

and

$$\operatorname{erf} = \int_0^x \exp(-y^2) dy$$

First we show that $\int_0^1 I_1(x) dx$ tends to zero as $n \rightarrow \infty$. Let a be a constant independent of x in the interval (0,1). For $0 \leq x \leq 1 - n^{-a}$ and all sufficiently large n we have

$$\begin{aligned}
 D &= \{(n-1)x^{2n+1} - nx^{2n-1} + x\}(1-x^2)^{-2} \\
 &= x(1-x^{2n})(1-x^2)^{-2} + 0\{n^{1+\alpha} \exp(-2n^{1-\alpha})\} \dots \dots \dots (2.9)
 \end{aligned}$$

and

$$B = (1+x^2)(1-x^{2n})(1-x^2)^{-3} + O\{n^{2+a} \exp(-2n^{1-a})\} \quad (2.10)$$

From (2.8) and (2.9) we can obtain

$$\Delta = (1-x^{2n})(1-x^2)^{-4} + O\{n^{2+2n} \exp(-2n^{1-a})\} \quad (2.11)$$

Now we choose $a=1-\{\log \log(n)^{10}\}/\log n$. Then since, for all sufficiently large

$$n, \quad n^{2+a} \exp(-2n^{1-a}) = n^{2+2n} \exp(-2 \log(n)^{10}) = n^{-18+a} \rightarrow 0.$$

All the error terms that appear in the formulas (2.9) to (2.11) will tend to zero.

Hence from (2.1), (2.9), (2.10), (2.11) and since for all x

$$(1-x^2)^{2/2} - x^2(1-x^2) + x^2(1+x^2)/2 > 1/5$$

we have

$$\begin{aligned} I_1(x) &= \int_0^{1-n^{-a}} (\Delta^{1/2} / wA) \exp[-m^2 \{(1-x^2)/(1-x^{2n})\}] \\ &\quad \times \left\{ (1-x^2)^2/2 - x^2(1-x^2) \right\} \\ &\quad + x^2(1+x^2)/2 \{1 + O\{n^{2+n} \exp(-2n^{1-a})\}\} dx \\ &\leq (1/\pi) \int_0^{1-n^{-a}} (1-x^2)^{-1} \exp\{-m^2(1-x^2)/5\} dx \\ &\leq (a/2\pi) \exp\{\log \log n \exp - (m^2/2n) \log(n)^{10}\} \dots \dots \quad (2.12) \end{aligned}$$

Now we note that since $m > \exp(n/\log n)$ the term m^2/n tends to infinity much faster than $\log \log n$ as $n \rightarrow \infty$, hence from (2.12) we can obtain

$$\int_0^{1-n^{-a}} I_1(x) dx \rightarrow 0 \text{ as } n \rightarrow \infty. \quad (2.13)$$

To show that $\int_0^{1-n^{-a}} I_1(x) dx \rightarrow 0$ as $n \rightarrow \infty$, we first prove that $(A - 2Dx + Bx^2)/\Delta$ is

positive for $1 - n^{-a} \leq x \leq 1$. For all sufficiently large n from (2.2) we have

$$\begin{aligned} A - 2Dx + Bx^2 &\geq Bx^2 - 2Dx \\ &\geq \{n^2 x^{2n} (1-x^2)^2 - 2nx^{2n+2} (1-x^2)\} \\ &\quad + x^2(1+x^2)(1-x^{2n}) \{1-x^2\} - 3 - n(n+1) \\ &\geq n^3 \{\log(n)^{10}\}^{-3} - 2n^2 > n^2 \quad (2.14) \end{aligned}$$

since $(1-x^2)^3 < (2n^{-a})^3$ and $2nx^{2n+2}(1-x^2) \rightarrow 0$ as $n \rightarrow \infty$.

Hence from (2.14) and since $\Delta < n^4$ we have

$$(A - 2Dx + Bx^2)/\Delta > n^{-2} \quad (2.15)$$

So from (2.15) and since from $(\Delta^2/a) < (2n-1)^{1/2}(1-x)^{-1/2}$,

we have
$$\int_0^{1-n^{-a}} I_1(x) dx < (2n-1)^{1/2} \exp(-m^2/n^2) \int_0^{1-n^{-a}} (1-x)^{-1/2} dx$$

$$< 3n^{(1-a)}/2 \exp(-m^2/n^2) \quad (2.16)$$

Which tends to zero as $n \rightarrow \infty$

In order to find $\int_0^{1-n^{-a}} I_1(x) dx$ we let $y = 1/x$, and divide the interval $0 \leq y \leq 1$ into three subintervals $(0, b)$, $(b, 1-(1/nd))$ and $(1-(1/nd), 1)$ where $d = \{(3/8) \log \log(n)\}^{1/3}$ and $b = (m^2 n d \log n)^{1/(4n-8)}$. We show that in these three subintervals

$$\int_0^{1-n^{-a}} I_1(x) dx = \int_0^1 y^{-2} I_1(1/y) dy = \log \{(1+b)/(1-b)\} \rightarrow 0 \text{ as } n \rightarrow \infty. \quad (2.17)$$

For $b \leq y \leq 1 - 1/nd$ from (2.2) we have

$$\{(A - 2D)/D\} + B/D^2 = 1 + \sum_{i=2}^n y^{-2i} + \sum_{i=2}^n y^{-2i} (i^2 - 2i) > n^3/4 \quad (2.18)$$

and

$$\Delta = \{1 - h^2(y)\} (1 - y^{2n}) / y^{4n-8} (1 - y^2) \leq n^2 / b^{4n-8} (1 - y^2)^2 < 3n^4 d^2 / b^{4n-8} \quad (2.19)$$

To prove that

$$\int_0^{\infty} I_1(x) dx \rightarrow 0 \text{ as } n \rightarrow \infty, \text{ from (2.2) for } y = 1/x \text{ we have .}$$

$$(Dx - A) / A^{1/2} \Delta^{1/2} = D^{2n-1} (1 - y^2)^{3/2} \{(n - 1 - ny - y^{2n-1})(1 - y^2)^2\} - y^{2n-4} \{1 - h^2(y)\}^{-1/2} (1 - y^{2n})^{-1/2} \quad (2.20)$$

And to find $\int_a^b I_2(x) dx$

$$\text{Let } x^2 / A = y^{2n-4} (1 - y^2) / (1 - y^{2n}) \quad (2.21)$$

First we let

$$0 \leq y \leq (mn^2)^{-1/(2n-1)} \text{ then } y \rightarrow 0 \text{ as } n \rightarrow \infty, \text{ from (2.21)}$$

$$\text{we have } (Dx - A) / A^{1/2} \Delta^{1/2} < 2ny^{2n-1} \quad (2.22)$$

$$\int_0^1 y^{-2} I_2(1/y) dy < \int_{m^2/2n^2}^{m/\sqrt{n}} \exp(-u^2/2) du$$

$$\leq (m/\sqrt{n} - m^2/2n^2) \exp(-m^2/4n^4) \rightarrow 0 \text{ as } n \rightarrow \infty. \quad (2.23)$$

Hence from (2.22) and (2.23) we have

$$\int_1^{\infty} I_2(x) dx \rightarrow 0 \text{ as } n \rightarrow \infty. \quad (2.24)$$

Finally from (2.8), (2.13), (2.16), (2.18) and (2.24) we obtain

$$EN(0, \infty) \leq 1/2$$

and since $EN(-\infty, \infty) = 2EN(0, \infty)$ we have proof of the theorem.

RESULT

The asymptotic number of crossings of the polynomial $P(x)$ with line $y=mx$ decreases as $m=0(\sqrt{n})$ increases. In this paper we proved that when $|m \geq \exp(nf)|$ the number of crossings reduces to one. The behaviour of the number of crossings between these two lines is not known. A subsequent study could consider the case when (m^2/n) tends to any non zero constant as n tends to infinity and as a guessed target $EN(-\infty, \infty) \sim (1/\pi) \log n$, which is half the number of crossings when $m=0$ seems reasonable. (Knowing a rough value for $EN(-\infty, \infty)$ is useful in order to sufficient upper and lower bounds for $EN(-\infty, \infty)$ leading to an asymptotic formula). Indeed, the behaviour of $EN(-\infty, \infty)$ for other values of m is also interesting, but it will involve more analysis especially for the $\int_{-\infty}^{\infty} I_2(x) dx$ part of $EN(-\infty, \infty)$. The result of the theorem shows that the expected number of crossings of a Random algebraic polynomial form $>\exp(f(n))$, where f is any function of n such that $f(n) \rightarrow \infty$, this expected number of crossings reduces to only one.

REFERENCES

1. H. Cramer and M.R. Lead better, *Stationary and Related Stochastic Process*. Wiley, New York, 2007.
2. I.A. Ibragimov and N.B. Maslova, *Soviet Math Dokl*, 2020; 12: 1004-1008.
3. K. Farhamand, *Stochastic Anal. Appl*, 2019; 6: 247-272.
4. M. Kac, *Bull Am. Math. Soc*, 2020; 46: 314-20.
5. Mishra, P.K. *Journal of Statistics and Mathematical Engineering*, 2022; 8: 1.
6. Mishra, P. K. & Dhal, D, *Theor. Prob. Appl*, 2021; 16: 228-48.
7. Mohanty, M.K. & Nayak, N. N, *Ann. Prob*, 2020; 14: 702-709.
8. Pratihary, D. & Muduli, J.C., *Indian J. Pure Appl. Math*, 2021; 20: 1-9.
9. S.O. Rice. *Bell System tech. J*, 2009; 25: 46-156.